

CS-GY 6923: Lecture 3

Regularization + Bayesian Perspective

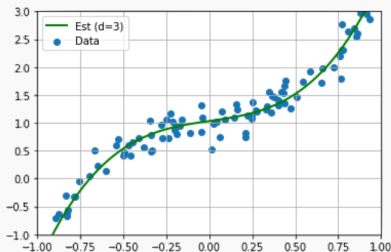
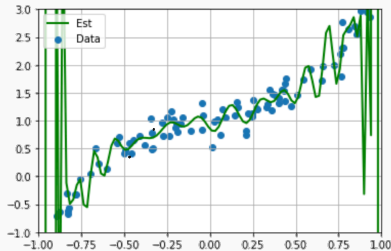
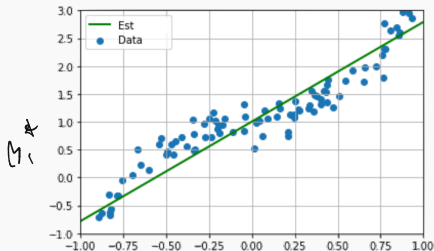
NYU Tandon School of Engineering, Prof. Christopher Musco

Model selection:

- Train models $f_{\theta_1}^{(1)}, \dots, f_{\theta_q}^{(q)}$ independently on training data to find optimal parameters $\theta_1^*, \dots, \theta_q^*$.
- Check loss $L_{test}(f_{\theta_1^*}^{(1)}), \dots, L_{test}(f_{\theta_q^*}^{(q)})$ on test data.
- Select mode with lowest test loss.

Can we used for arbitrary sets of models. Often used when you are not sure how “complex” your model should be for the data, and want to find the sweet spot between a good fit, and not overfitting.

LAST CLASS + LAB



\mathcal{L}^1

Underfit, overfit, just right.

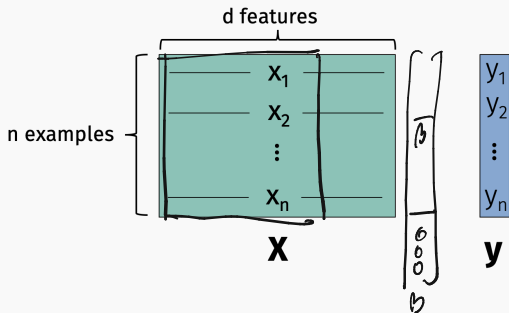
In the model selection examples we discussed last class, we had full control over the complexity of the model: could range from underfitting to overfitting.

In practice, we often don't have this freedom. Even the most basic model might lead to overfitting.

OVER-PARAMETERIZED MODELS

$$\bar{X} \bar{\beta} = \bar{y} \quad \bar{\beta} = \bar{X}^{-1} \bar{y}$$

Example: Linear regression model where $d \geq n$.



Can (almost) always find β so that $X\beta$ = y exactly.

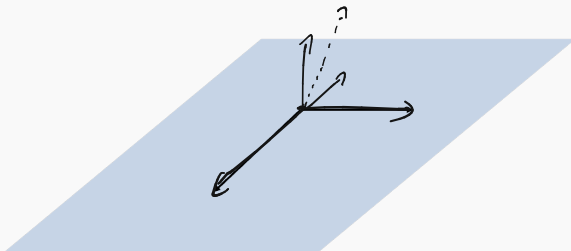
(condition when this true:

As long as there are n linearly independent features.

HIGH DIMENSIONAL LINEAR MODELS

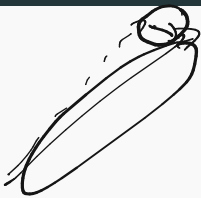
Claim: For almost all sets of n , length n vectors $\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(n)}$, we can write any vector \mathbf{y} as a linear combination of these vectors.

$n=3$



I.e., can find some coefficients so that
 $\beta_1 \mathbf{x}^{(1)} + \dots + \beta_q \mathbf{x}^{(q)} = \mathbf{X}\boldsymbol{\beta} = \mathbf{y}.$

ZERO TRAIN LOSS

$$\begin{bmatrix} 1 \\ x_1 \\ \vdots \\ x_n \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_n \end{bmatrix} \quad \underbrace{\beta_0}_{y_{10}} x^{(1)} + \underbrace{\beta_1}_{x^{(2)}} x^{(2)} \approx y$$


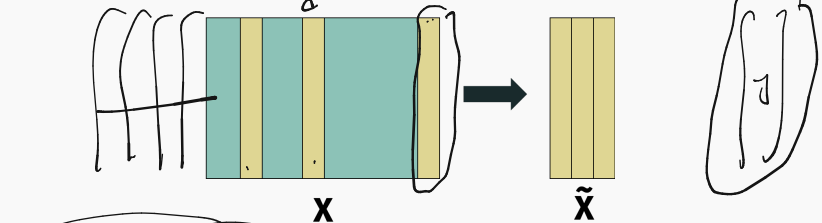
- We will discuss some models later in the class where zero training loss is not necessarily a bad sign: k -nearest neighbors, some neural nets.
- Typically however it will be a sign of overfitting, as in the polynomial regression example.

$$R^2 = 1 - \frac{\|x_n - y\|_2^2}{\|y\|_2^2} \quad \frac{\|y\|_2^2}{\|y\|_2^2}$$

$$= 0 \quad = 1$$

FEATURE SELECTION

Select some subset of $\ll n$ features to use in model:



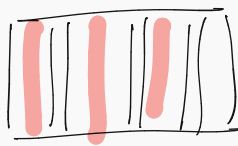
Filter method: Compute some metric for each feature, and select features with highest score.

- Example: compute loss or R^2 value when each feature in X is used in single variate regression.

(Any potential limitations of this approach?)

FEATURE SELECTION

$$q = 2 \dots 10$$

$$\frac{(d)(d-1)\dots(d-q+1)}{q!} \approx d^q$$


Exhaustive approach: Pick best subset of q features. Train $\binom{d}{q}$ models.



$$\begin{bmatrix} 1 \\ 0 \\ 1 \\ 1 \\ 0 \\ 1 \end{bmatrix}$$

$$q > b$$

$$B =$$

$$\begin{bmatrix} 100 \\ 10 \\ .5 \\ 100 \end{bmatrix}$$

FEATURE SELECTION


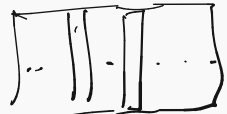


Diagram showing a single feature selection step. A box is divided into two parts: a shaded part on the left and an unshaded part on the right. Above the box, the expression $d + d-1 + \dots + 1$ is written, with a note $\approx 100 - 90$ and $O(d)$ below it.

Faster approach: Greedily select q features.



Stepwise Regression:

$$\underline{d} + \underline{d-1} + \dots + \underline{1} \approx \textcircled{qd}$$

- Forward:** Step 1: pick single feature that gives lowest loss.

Step k : pick feature that when combined with previous $k - 1$ chosen features gives lowest loss.

- Backward:** Start with all of the features. Greedily eliminate those which have least impact on model performance.

Feature selection deserves more than two slides, but we won't go into too much more detail!

$$d \cdot (d-1) \approx O(d^2)$$



$d=3$ 4-variable regression problem ^{regression} ¹⁰

ALTERNATIVE APPROACH

Regularization: Discourage overfitting by adding a regularization penalty to the loss minimization problem.

$$\beta^* = \underset{\beta}{\operatorname{arg\,min}} L(\beta)$$

$$\min_{\beta} [L(\beta) + \textcolor{brown}{Reg}(\beta)].$$

positive constant

Example: Least squares regression. $L(\beta) = \|X\beta - y\|_2^2$.

- Ridge regression (ℓ_2): $\textcolor{brown}{Reg}(\beta) = \lambda \|\beta\|_2^2$
- LASSO (least absolute shrinkage and selection operator) (ℓ_1): $\textcolor{brown}{Reg}(\beta) = \lambda \|\beta\|_1 = \sum_{i=1}^p |\beta_i|$
- Elastic net: $\textcolor{brown}{Reg}(\beta) = \lambda_1 \|\beta\|_1 + \lambda_2 \|\beta\|_2^2$

Note that $\arg \min_{\beta} [L(\beta) + \textcolor{brown}{Reg}(\beta)] \neq \arg \min_{\beta} [L(\beta)]$

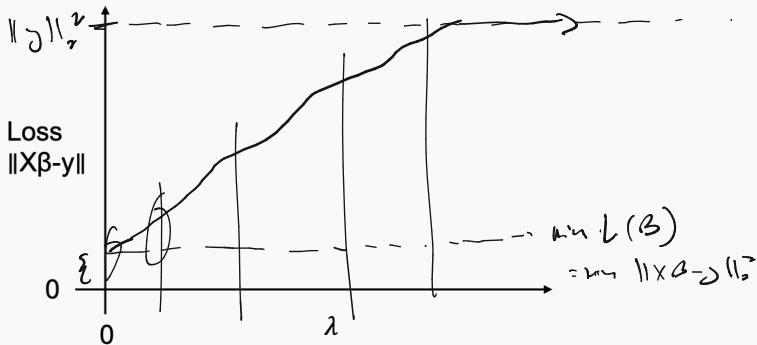
RIDGE REGULARIZATION: PERSPECTIVE 1

$\frac{1}{2} \times \frac{1}{2} = \frac{1}{4}$

$$XB \rightarrow O$$

Ridge regression: $\min_{\beta} (\|X\beta - y\|_2^2) + (\lambda \|\beta\|_2^2)$.

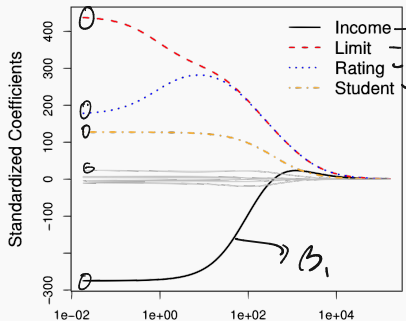
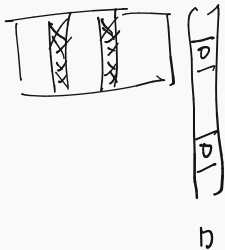
- As $\lambda \rightarrow \infty$, we expect $\|\beta\|_2^2 \rightarrow 0$ and $\|X\beta - y\|_2^2 \rightarrow \|y\|_2^2$.
- By choosing different values of λ we have models of varying levels of model fit.



RIDGE REGULARIZATION

$$\text{Ridge regression: } \min_{\beta} \|X\beta - y\|_2^2 + \lambda \|\beta\|_2^2.$$

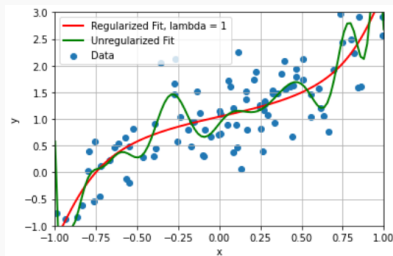
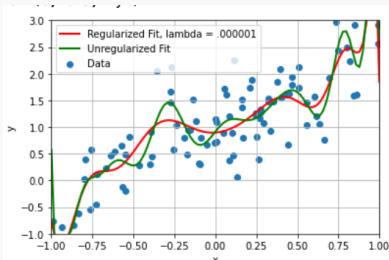
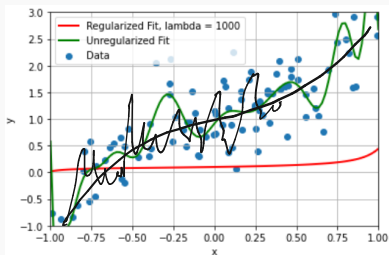
- As $\lambda \rightarrow \infty$, we expect $\|\beta\|_2^2 \rightarrow 0$ and $\|X\beta - y\|_2^2 \rightarrow \|y\|_2^2$.
- Feature selection attempts to set many coordinates in β to 0. Regularization encourages coordinates to be small.



(Can be viewed as a “soft” version of feature selection.)

POLYNOMIAL EXAMPLES

Fit degree 20 polynomial with varying levels of regularization.



RIDGE REGULARIZATION

How do we minimize: $L_R(\beta) = \underbrace{\|X\beta - y\|_2^2}_{f(\beta)} + \underbrace{\lambda\|\beta\|_2^2}_{g(\beta)}$?

$$\underline{\nabla L_R(\beta)} = 0$$

$$\nabla(f+g) = \nabla f + \nabla g$$

$$\nabla \|X\beta - y\|_2^2 + \nabla (\lambda\|\beta\|_2^2)$$

$$(\text{dxd} \text{Id} + \text{Id} + \text{Id})$$

$$2X^T(X\beta - y) + 2\lambda\beta = 0$$

$$X^T X \beta - X^T y + \lambda I \beta$$

$$= (X^T X + \lambda I) \beta - X^T y = 0$$

$$(X^T X + \lambda I) \beta = X^T y$$

$$\beta^* = (X^T X + \lambda I)^{-1} X^T y$$

RIDGE REGULARIZATION

How do we minimize: $L_R(\beta) = \|X\beta - y\|_2^2 + \lambda\|\beta\|_2^2$?

$$\beta^* = (X^T X + \lambda I)^{-1} \underline{X^T y}$$

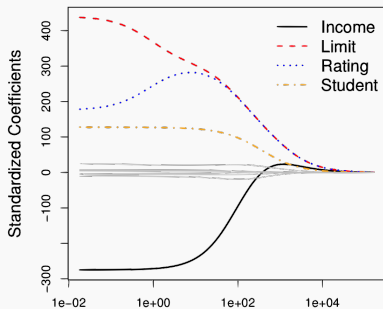
$$\frac{1}{\lambda} X^T y$$

The diagram shows a square matrix on the left with many small λ terms scattered throughout. An arrow points to the right, where a large λ is written before another square matrix. This second matrix has a diagonal of ones (represented by vertical lines) and zeros elsewhere, representing the identity matrix I in the expression $(X^T X + \lambda I)^{-1}$.

LASSO REGULARIZATION

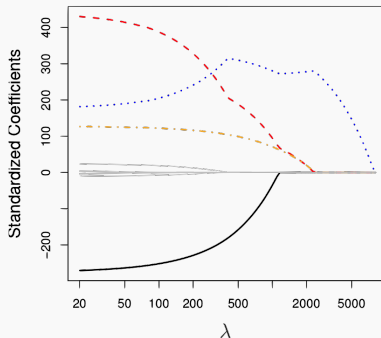
Lasso regularization: $\min_{\beta} \|X\beta - y\|_2^2 + \lambda \|\beta\|_1$

- As $\lambda \rightarrow \infty$, we expect $\|\beta\|_1 \rightarrow 0$ and $\|X\beta - y\|_2^2 \rightarrow \|y\|_2^2$.
- Typically encourages subset of β_i 's to go to zero, in contrast to ridge regularization.



1 2 ... 20

Ridge regularization



Lasso Regularization

Why Lasso encourages sparsity is a long story that was only understand relatively recently. Major topic in the field of (compressed sensing) and (sparse recovery.)

Pros:

- Simpler, more interpretable model.
- More intuitive reduction in model order.

Cons:

- No closed form solution because $\|\beta\|_1$ is not differentiable.
- Can be solved with iterative methods, but generally not as quickly as ridge regression.

REGULARIZATION

$$\lambda = 10, 20, 30, \dots$$

$$\frac{1}{16} \quad \frac{1}{8} \quad \frac{1}{4} \quad \dots \quad 16$$

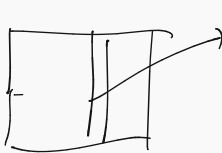
Notes:

Model selection/cross validation used to choose optimal scaling λ on $\lambda \|\beta\|_2^2$ or $\lambda \|\beta\|_1$.

Often grid search for best parameters is performed in “log space”. E.g. consider $[\lambda_1, \dots, \lambda_q] = \frac{1}{2^k} [-4, -3, -2, -1, -0, 1, 2, 3, 4]$.

Regularization methods are not invariant to data scaling.

Typically when using regularization we mean center and scale columns to have unit variance.


$$\frac{(x_i - \bar{x}_i)}{\|x_i - \bar{x}_i\|_2}$$

Break Until
3:25.

THE BAYESIAN/PROBABILISTIC MODELING PERSPECTIVE

CLASSIFICATION SETUP

- **Data Examples:** $\mathbf{x}_1, \dots, \mathbf{x}_n \in \mathbb{R}^d$
- **Target:** $y_1, \dots, y_n \in \{0, 1, \dots, q-1\}$ when there are q classes.
 - Binary Classification: $q = 2$, so each $y_i \in \{0, 1\}$.
 - Multi-class Classification: $q > 2$.¹

¹Note that there is also multi-label classification where each data example may belong to more than one class.

CLASSIFICATION EXAMPLES

- Medical diagnosis from MRI: 2 classes.
- MNIST digits: 10 classes.
- Full Optical Character Recognition: 100s of classes.
- ImageNet challenge: 21,000 classes.

Running example today: Email Spam Classification.

Classification can (and often is) solved using the same **loss-minimization framework** we saw for regression.

We won't see that today! We're going to use classification as a window into another way of thinking about machine learning.

Will give new, interesting justifications for tools like regularization. will also lead to natural approaches for generative ML.

Rest of Today: ML from a **Probabilistic Modeling/Bayesian Perspective**.

In a Bayesian or Probabilistic approach to machine learning we always start by conjecturing a

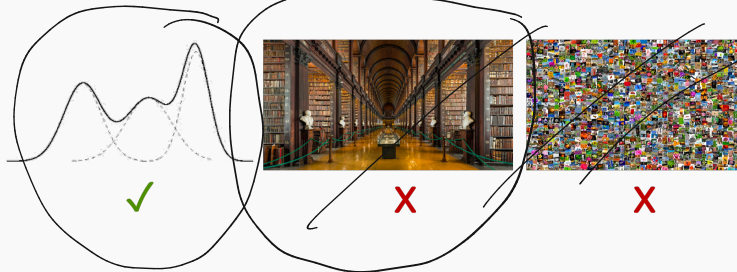
probabilistic model

that plausibly could have generated our data.²

- The model guides how we make predictions.
- The model typically has unknown parameters $\vec{\theta}$ and we try to find the most reasonable parameters based on observed data (more on this later in lecture).

²“Data” here includes both the predictors x_1, \dots, x_n and targets y_1, \dots, y_n .

Typically we try to keep things simple!



Exercise: Come up with a probabilistic model for the following data set $(x_1, y_1), \dots, (x_n, y_n)$.

- For n **NYC apartments**: each x_i is the size of the apartment in square feet. Each y_i is the monthly rent in dollars.

What are the unknown parameters of your model. What would be a guess for their values? How would you confirm or refine this guess using data?

PROBABILISTIC MODELING

Dataset: $(x_1, y_1), \dots, (x_n, y_n)$



Description: For n NYC apartments: each x_i is the size of the apartment in square feet. Each y_i is the monthly rent in dollars.

Model:

$$x_i \sim \mathcal{N}(1000, 6^2) \quad \text{option 1}$$

$$x_i \sim \text{Unif}(300, \dots, 10000) \quad \text{option 2}$$

$$y_i = \underbrace{\beta_1}_{\text{weight}} x_i + \underbrace{\beta_0}_{\text{bias}} + \mathcal{N}(0, 6^2)$$

PROBABILISTIC MODELING

Dataset: $(x_1, y_1), \dots, (x_n, y_n)$

Description: For n undergraduate **students**: each $x_i \in \{1, 2, 3, 4\}$ indicating class year. Each $y_i \in \{0, 1\}$ with zero indicating the student has not taken machine learning, one indicating they have.

Model:

$$x_i \sim \text{Unif}(1, 2, 3, 4)$$

$$y_i \sim \text{Ber}(p(x_i))$$

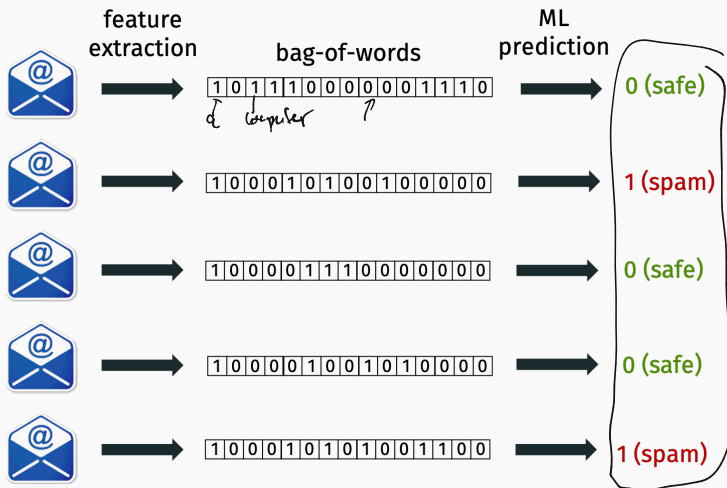
0 1
w/o w/p

$p(1) = .1 \quad p(2) = .2 \quad p(3) = .7 \quad p(4) = .99$

Goal:

- Build a probabilistic model for a binary classification problem.
- Estimate parameters of the model. }
- From the model derive a classification rule for future predictions (the Naive Bayes Classifier).

SPAM PREDICTION



Both target labels and data vectors are binary.

EMAIL MODEL

$$[\underset{P_1}{1} \underset{P_2}{0} \underset{P_3}{0} \underset{P_4}{0} \underset{P_5}{1} \underset{P_6}{0} \underset{P_7}{0} \underset{P_8}{0} \underset{P_9}{0} \underset{P_{10}}{0} \underset{P_{11}}{0} \underset{P_{12}}{0}]$$

Let's create a probabilistic model that generates emails.

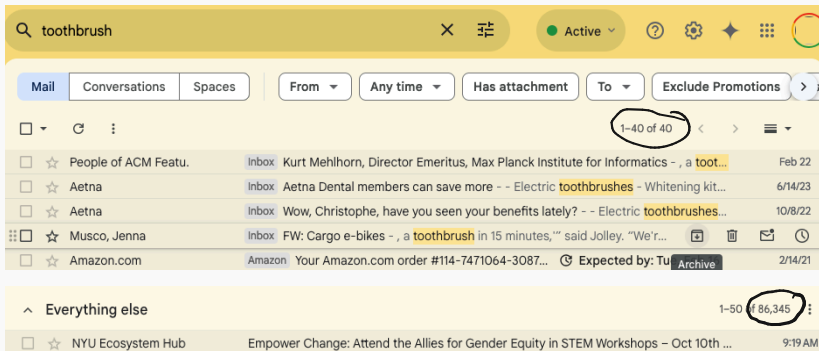
Observation: Since bag-of-words features don't care about word order, our model does not need to either.

- Common approach. Assign a probability $p_i \in [0, 1]$ to word i . Set $\mathbf{x}_i = 1$ with probability p_i , $\mathbf{x}_i = 0$ with probability $1 - p_i$.

$$p_{\text{the}} = .9 \quad p_{\text{calendar}} = .2 \quad \underline{p_{\text{toothbrush}} = .01}$$

EMAIL MODEL

Model training: Find parameters p_1, \dots, p_d that best fit our training data.



In this case, set p_i to empirical word frequency of word i .

$$p_{\text{toothbrush}} = \frac{40}{86345}$$

How can we make this model richer to generate both spam and non-spam email?

- Different words tend to be more or less frequent in spam or regular emails.

Not Spam

$$p_{\text{won}} = .01$$

$$p_{\text{\$}} = .1$$

$$p_{\text{student}} = .1$$

Spam

$$p_{\text{won}} = .1$$

$$p_{\text{\$}} = .2$$

$$p_{\text{student}} = .01$$

PROBABILISTIC MODEL FOR EMAIL

Probabilistic model for (bag-of-words, label) pair (\mathbf{x}, y) :

$c_0 = 95\%$

- Set $y = 0$ with probability $\underline{c_0}$, $y = 1$ with probability $\underline{c_1 = 1 - c_0}$.
 - c_0 is probability an email is not spam (e.g. 99%).
 - c_1 is probability an email is spam (e.g. 1%).
- If $y = 0$ for each i , set $x_i = 1$ with prob. $p_i^{(0)}$.
- If $y = 1$ for each i , set $x_i = 1$ with prob. $p_i^{(1)}$.

X

Unknown model parameters:

- c_0, c_1 ,
- $(p_1^{(0)}, p_2^{(0)}, \dots, p_d^{(0)})$ \rightarrow non-spam prob. one for each of the d vocabulary words.
- $p_1^{(1)}, p_2^{(1)}, \dots, p_d^{(1)}$ \rightarrow spam prob. one for each of the d vocabulary words.

How would you estimate these parameters?

Reasonable way to set parameters:

- Set $\overleftarrow{c_0}$ and c_1 to the empirical fraction of not spam/spam emails. —
- { For each word i , set $p_i^{(0)}$ to the empirical probability word i appears in a non-spam email.
- { For each word i , set $p_i^{(1)}$ to the empirical probability word i appears in a spam email.

DONE WITH MODELING
ON TO PREDICTION

PROBABILITY REVIEW

$$p(A \cap B)$$

word

- **Probability:** $p(A)$ – the probability event A happens.
- **Joint probability:** $p(A, B)$ – the probability that event A and event B happen.
- **Conditional Probability** $p(A | B)$ – the probability A happens given that B happens.

$$p(A | B) = \frac{p(A, B)}{p(B)}$$
$$\underline{p(A | B)} \underline{p(B)} = p(A, B)$$

$$\underline{P_f = P_r = P}$$

Two random events are independent if:

$$\Pr(\underline{A} \mid \underline{B}) = \Pr(A), \quad \text{or equivalently,} \quad \Pr(\underline{B} \mid \underline{A}) = \Pr(B)$$

Equivalent characterization:

$$\underline{\Pr(A, B)} = P(A) \cdot P(B).$$

$$\frac{P(A/B)}{P(A)} = \frac{P(A, B)}{P(B)}$$

$$P(A)P(B) = P(A, B)$$

PROBABILITY REVIEW

B : event that dice roll is even.

Note that when we write something like $p(A | B)$, A and B are random events not random variables.

We will sometimes (informally) write $p(X | B)$, where X is a random variable. In this case, $p(X | B)$ is understood to be a probability density/mass function.

E.g., suppose X is a dice roll that takes values $\underline{1}, \dots, \underline{6}$. Then $p(X | B)$ is a function from $\{1, \dots, 6\} \rightarrow [0, 1]$ whose i^{th} value equals $\underline{p(X = i | B)}$. $A = \{X = i\}$ is a proper random event.

$$p(X | B) = \begin{array}{cccccc} 0 & 1/3 & 0 & 1/3 & 0 & 1/3 \\ 1 & 2 & 3 & 4 & 5 & 6 \end{array} \quad p(X=1 | B)$$

BAYES THEOREM/RULE

$$\left(P(A|B) = \frac{P(B|A) P(A)}{P(B)} \right)$$

~~$$P(A|B) = \frac{P(A|B) P(A)}{P(B)}$$~~

Proof:

$$P(A|B) = \frac{P(A, B)}{P(B)}$$

$$P(B|A) = \frac{P(A, B)}{P(A)}$$

$$P(B) P(A|B) = P(A) P(B|A)$$

$$\left(P(A|B) = \frac{P(A) P(B|A)}{P(B)} \right)$$

Bayes Rule.

CLASSIFICATION RULE

Given unlabeled input ($\underline{\mathbf{w}}$, ---), choose the label $y \in \{0, 1\}$ which is most likely given the data. Recall $\mathbf{w} = [0, 0, 1, \dots, 1, 0]$.

Classification rule: **maximum a posterior (MAP) estimate**.

Step 1. Compute: $p_r(y=0 | \mathbf{w} = [0, 0, 1, \dots, 1, 0])$

- $\underline{p(y = 0 | \mathbf{w})}$: prob. $y = 0$ given observed data vector \mathbf{w} .
- $\underline{p(y = 1 | \mathbf{w})}$: prob. $y = 1$ given observed data vector \mathbf{w} .

Step 2. Output: 0 or 1 depending on which probability is larger.

$p(y = 0 | \mathbf{w})$ and $p(y = 1 | \mathbf{w})$ are called posterior probabilities.

EVALUATING THE POSTERIOR

How to compute the posterior? **Bayes rule!**

$$p(y=0 | \mathbf{w}) = \frac{p(\mathbf{w} | y=0)p(y=0)}{\underline{\underline{p(\mathbf{w})}}}$$

(1)

Handwritten notes: An arrow points from the text "Bayes rule!" to the equation. The variable y in the denominator is circled. The expression $p(\mathbf{w})$ in the denominator is underlined twice. Above the equation, the handwritten text $Pr(\mathbf{w} = [001000] | y=0)$ is written.

$$\text{posterior} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}$$

(2)

Handwritten notes: The word "posterior" is circled. The word "evidence" is circled.

- **Prior:** Probability in class 0 prior to seeing any data.
- **Posterior:** Probability in class 0 after seeing the data.

EVALUATING THE POSTERIOR

Goal is to determine which is larger:

$$\begin{aligned} p(y = 0 | \mathbf{w}) &= \frac{p(\mathbf{w} | y = 0)p(y = 0)}{p(\mathbf{w})} \\ p(y = 1 | \mathbf{w}) &= \frac{p(\mathbf{w} | y = 1)p(y = 1)}{p(\mathbf{w})} \end{aligned} \quad \text{vs.}$$

- We can ignore the evidence $p(\mathbf{w})$ since it is the same for both sides!
- $p(y = 0)$ and $p(y = 1)$ already known (computed from training data). These are our computed parameters $\underline{c_0}, \underline{c_1}$.
5% 5%
- $p(\mathbf{w} | y = 0)$ = ? $p(\mathbf{w} | y = 1)$ = ?

EVALUATING THE POSTERIOR

Consider the example ($\mathbf{w} = [0, 1, 1, 0, 0, 0, 1, 0]$.)

Recall that, under our model, index i is 1 with probability $\underline{p_i^{(0)}}$ if we are not spam, and 1 with probability $p_i^{(1)}$ if we are spam.

$$p(\mathbf{w} = [0, 1, 1, 0, 0, 0, 1, 0] \mid y = 0) =$$

$$(1 - p_1^{(0)}) \underbrace{p_2^{(0)}} p_3^{(0)} (1 - p_4^{(0)}) \dots (1 - p_m^{(0)})$$

$$p(\mathbf{w} = [0, 1, 1, 0, 0, 0, 1, 0] \mid y = 1) =$$

$$(1 - p_1^{(1)}) p_2^{(1)} p_3^{(1)} \dots$$

Final Naive Bayes Classifier

Training/Modeling: Use existing data to compute:

- Prior class probabilities $c_0 = p(y = 0)$, $c_1 = p(y = 1)$
- For all i :
 - $p_i^{(0)} = p(w_i = 1 | y = 0)$ and $(1 - p_i^{(0)}) = p(w_i = 0 | y = 0)$
 - $p_i^{(1)} = p(w_i = 1 | y = 1)$ and $(1 - p_i^{(1)}) = p(w_i = 0 | y = 1)$

Prediction:

- For new input $\mathbf{w} \in \{0, 1\}^d$:
 - Compute $p(\mathbf{w} | y = 0) = \prod_{i=1}^d p(w_i | y = 0)$
 - Compute $p(\mathbf{w} | y = 1) = \prod_{i=1}^d p(w_i | y = 1)$
- Return

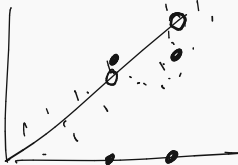
$$\arg \max [\underbrace{p(\mathbf{w} | y = 0)}_{.45} \cdot \underbrace{p(y = 0)}_{.95}, \underbrace{p(\mathbf{w} | y = 1)}_{.05} \cdot \underbrace{p(y = 1)}_{.05}] .$$

OTHER APPLICATIONS OF
THE BAYESIAN PERSPECTIVE

BAYESIAN REGRESSION

The Bayesian view offers an interesting alternative perspective on many machine learning techniques.

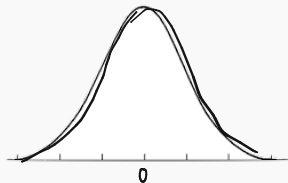
Example: Linear Regression



Probabilistic model: For some “true” set of parameters β_{true} ,

$$y = \langle \mathbf{x}, \beta_{\text{true}} \rangle + \eta$$

where the η drawn from $N(0, \sigma^2)$ is random Gaussian noise.



$$Pr(\eta = z) \sim \underline{\underline{e^{-z^2}}}$$

The symbol \sim means “is proportional to”.

GAUSSIAN DISTRIBUTION REFRESHER

Names for same thing: Normal distribution, Gaussian distribution, bell curve.

Parameterized by mean μ and variance σ^2 .



η is a continuous random variable, so it has a probability density function $p(\eta)$ with $\int_{-\infty}^{\infty} p(\eta) d\eta = 1$

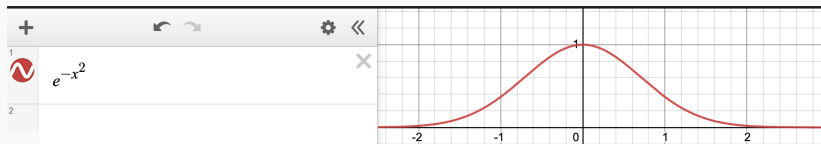
$$p(\eta) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{\eta-\mu}{\sigma}\right)^2}$$

$$e^{-\frac{\eta^2}{2\sigma^2}}$$

GAUSSIAN DISTRIBUTION REFRESHER

$$e^{-x^2}$$

The important thing to remember is that the the PDF falls off exponentially as we move further from the mean.



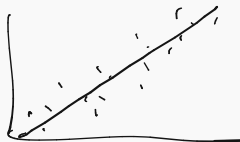
The normalizing constant in front $1/2$, etc. don't matter so much.

BAYESIAN REGRESSION

Example: Linear Regression.

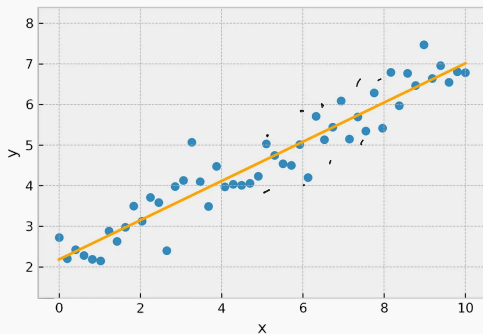
Probabilistic model:

$$y = \langle \mathbf{x}, \underline{\beta_{\text{true}}} \rangle + \eta$$



where the η drawn from $N(0, \sigma^2)$ is **random Gaussian noise**.

The noise is independent for different inputs $\mathbf{x}_1, \dots, \mathbf{x}_n$.



How should we find the unknown parameters β for our model?

$$p(x, y | \beta)$$

Also use a Bayesian approach!

First thought: choose β to maximize:

$$\text{posterior} = \Pr(\beta | \underline{X, y}) = \frac{\Pr(X, y | \beta) \Pr(\beta)}{\Pr(X, y)} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}.$$

But in this case, we don't have a prior – no values of β are inherently more likely than others.

Choose β to maximize just the likelihood:

$$\frac{\Pr(X, y | \beta) \cancel{\Pr(\beta)}}{\Pr(X, y)} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}.$$

This is called the **maximum likelihood estimate**.

FIXED DESIGN LINEAR REGRESSION

$$\Pr(x, y | \beta) = \Pr(y | \beta)$$

Often we think of \mathbf{X} as fixed and deterministic, and only \mathbf{y} is generated at random in the model. This is called the fixed design setting. Can also consider a randomized design setting, but it is slightly more complicated.

In the fixed design setting our task of maximizing $\Pr(\mathbf{X}, \mathbf{y} | \beta)$ simplifies to maximizing

$$\max_{\beta} \Pr(\mathbf{y} | \beta)$$

MAXIMUM LIKELIHOOD ESTIMATE

Data:

$$X = \begin{bmatrix} - & x_1 & - \\ - & x_2 & - \\ & \vdots & \\ - & x_n & - \end{bmatrix} \quad y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

Model: $y_i = \langle x_i, \beta \rangle + \eta_i$ where $p(\eta_i = z) \sim \frac{e^{-z^2/2\sigma^2}}{\sigma\sqrt{2\pi}}$ and η_1, \dots, η_n are independent.

$$\begin{aligned} \Pr(y | \beta) &\sim \prod_{i=1}^n \Pr(y_i | \beta) = \prod_{i=1}^n \Pr(\eta_i = y_i - \langle x_i, \beta \rangle) \\ &= \prod_{i=1}^n e^{-(y_i - \langle x_i, \beta \rangle)^2 / 2\sigma^2} \end{aligned}$$

Easier to work with the **log likelihood**:

$$\begin{aligned}
 \arg \max_{\beta} \Pr(\mathbf{y} \mid \beta) &= \arg \max_{\beta} \prod_{i=1}^n e^{-(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2} \\
 &= \arg \max_{\beta} \log \left(\prod_{i=1}^n e^{-(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2} \right) \\
 &= \arg \max_{\beta} \sum_{i=1}^n \log \left(e^{-(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2} \right) \\
 &= \arg \max_{\beta} \sum_{i=1}^n -(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2 \\
 &= \arg \min_{\beta} \sum_{i=1}^n (y_i - \langle \mathbf{x}_i, \beta \rangle)^2.
 \end{aligned}$$

Conclusion: Choose β to minimize:

$$\left(\sum_{i=1}^n (y_i - \langle \mathbf{x}_i, \beta \rangle)^2 = \underbrace{\|\mathbf{y} - \mathbf{X}\beta\|_2^2} \right)$$

This is a completely different justification for minimizing squared loss!

Minimizing the ℓ_2 loss is “optimal” when you assume your data follows a linear model with i.i.d. Gaussian noise (with any fixed variance).