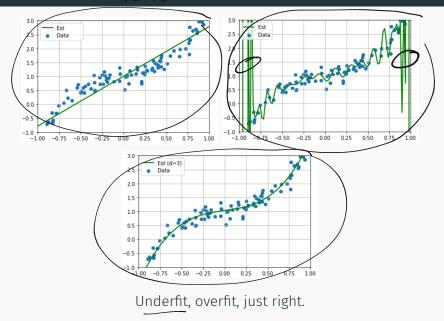
CS-GY 6923: Lecture 3 Regularization + Bayesian Perspective

NYU Tandon School of Engineering, Prof. Christopher Musco

- Model selection: $f_{\theta_1}^{(r)}$, $f_{\theta_q}^{(r)}$ Train models $M_{\theta_1}, \dots, M_{\theta_q}$ independently on training data to find optimal parameters $\theta_1^*, \ldots, \theta_a^*$.
 - Check loss $L_{test}(f_1, \theta_1^*), \ldots, L_{test}(f_a, \theta_a^*)$ on test data.
 - Select mode with lowest test lost.

Can we used for arbitrary sets of models. Often used when you are not sure how "complex" your model should be for the data, and want to find the sweet spot between a good fit, and not overfitting.

LAST CLASS + 🕼

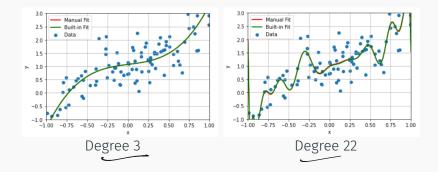


In the lab we had you use **numpy.polynomial.polynomial**. polynomial. Last class, however, we discussed how we could use multiple linear regression instead. If our point to fit at are $x_1, \ldots, x_n \in [-1, 1]$, would construct the data matrix:

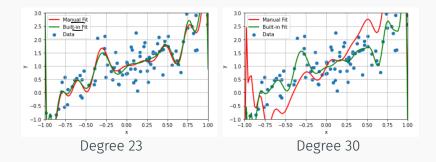
$$\mathbf{X} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ 1 & x_3 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \begin{pmatrix} x_1^2 & x_1^3 \\ x_2^2 & x_2^3 \\ x_3^2 & x_3^3 \\ \vdots & \vdots \\ x_n^2 & x_n^3 \end{bmatrix}$$

Then find polynomial coefficents as $\underline{\beta} = (X^T X)^{-1} X^T y$.

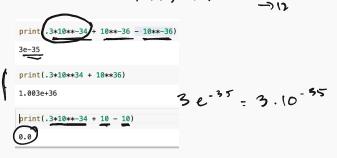


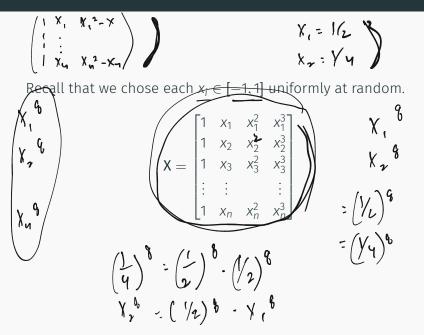


```
# built in function
beta_hat = poly.polyfit(xdat,ydat,d)
# manual fit using naive multivariate regression
X = np.zeros([len(xdat),d+1])
for i in range(d+1):
X[:,i] = xdat**i
my_beta = np.linalg.inv(np.transpose(X)@X)@np.transpose(X)@ydat
```



- Your computer can easily deal with both very large and very small numbers. Underflow and overflow are extremely unlikely to be issues in floating point arithmetic.
- The issue is when you compute using numbers of very differing magnitude.
 4%5 . 10⁻⁰⁰⁰ 6





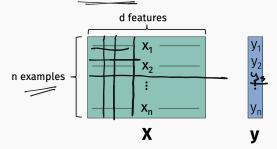
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REGULARIZATION

In the model selection examples we discussed last class, we had full control over the complexity of the model: could range from <u>underfitting</u> to <u>overfitting</u>.

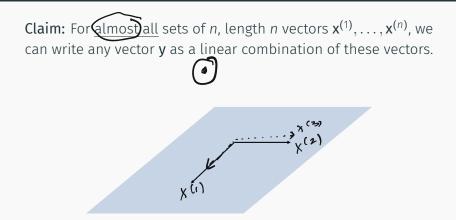
In practice, you often don't have this freedom. Even the <u>most</u> <u>basic model</u> might lead to overfitting.

Example: Linear regression model where $d \ge n$.



Can (almost) always find β so that $X\beta = y$ exactly. $X\beta = \beta, \chi^{(1)} + \beta, \chi^{(2)} + \beta, \chi^{(3)} = \beta$

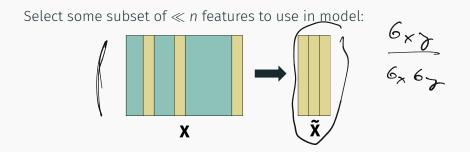
HIGH DIMENSIONAL LINEAR MODELS



I.e., can find some coefficients so that $\beta_1 \mathbf{x}^{(1)} + \ldots + \beta_q \mathbf{x}^{(q)} = \mathbf{X}\boldsymbol{\beta} = \mathbf{y}.$

- We will discuss some models later in the class where zero training loss is not necessarily a bad sign: *k*-nearest neighbors, some neural nets.
- Typically however if will be a sign of overfitting, as in the polynomial regression example.

FEATURE SELECTION

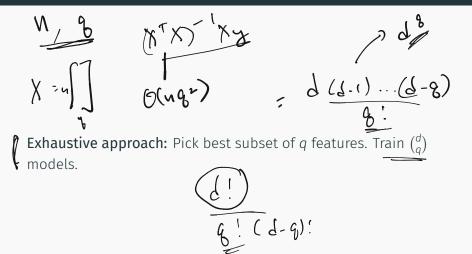


Filter method: Compute some metric for each feature, and select features with highest score.

• Example: compute loss or *R*² value when each feature in **X** is used in single variate regression.

Any potential limitations of this approach?

FEATURE SELECTION



FEATURE SELECTION

Faster approach: Greedily select q features.

Stepwise Regression:

- $d = \frac{d}{d-1} = \frac{d}{d-2} = \frac{d}{d-8}$ dily select q features. $d = \frac{d}{d-8}$
- Forward: Step 1: pick single feature that gives lowest loss.
 Step k: pick feature that when combined with previous k 1 chosen features gives lowest loss.
- 🕐 Backward: Start with all of the features. Greedily eliminate 🏌
- those which have least impact on model performance.

Feature selection deserves more than two slides, but we won't go into too much more detail! $d = 4 \cdot 5$

ALTERNATIVE APPROACH

Regularization: Discourage overfitting by adding a regularization penalty to the loss minimization problem.

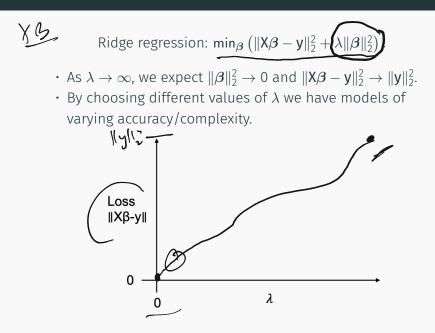
$$\min_{\beta} \left[L(\beta) + \frac{\operatorname{Reg}(\beta)}{\beta} \right]. \qquad \qquad \bigwedge > \circlearrowright$$

Example: Least squares regression. $L(\beta) = ||X\beta - y||_2^2$.

- Ridge regression (ℓ_2) : $Reg(\beta) = \lambda \|\beta\|_2^2 = \frac{1}{2} \beta_1^2$
- LASSO (least absolute shrinkage and selection operator) $(\ell_1): \operatorname{Reg}(\beta) = \lambda \|\beta\|_1 \rightarrow \xi |\beta|_1$
- Elastic net: $Reg(\beta) = \lambda_1 \|\beta\|_1 + \lambda_2 \|\beta\|_2^2$

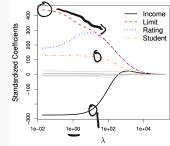
Note that $\operatorname{arg\,min}_{\beta}[L(\beta) + \operatorname{Reg}(\beta)] \neq \operatorname{arg\,min}_{\beta}[L(\beta)]$

RIDGE REGULARIZATION: PERSPECTIVE 1



Ridge regression: $\min_{\beta} \left(\|\mathbf{X}\beta - \mathbf{y}\|_2^2 + \lambda \|\beta\|_2^2 \right)$.

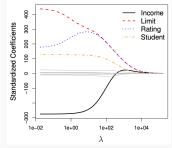
- As $\lambda \to \infty$, we expect $\|\beta\|_2^2 \to 0$ and $\|\mathbf{X}\beta \mathbf{y}\|_2^2 \to \|\mathbf{y}\|_2^2$.
- Feature selection methods attempt to set many coordinates in β to 0. Ridge regularizations encourages coordinates to be close to zero.





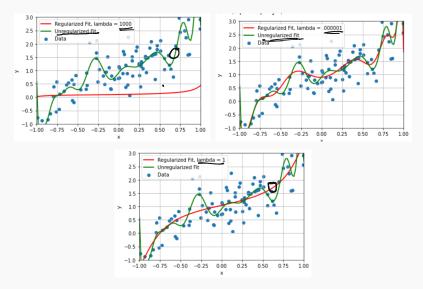
Ridge regression:
$$\min_{\beta} \|\mathbf{X}\beta - \mathbf{y}\|_{2}^{2} + \lambda \|\beta\|_{2}^{2}$$
.

- As $\lambda \to \infty$, we expect $\|\beta\|_2^2 \to 0$ and $\|\mathbf{X}\beta \mathbf{y}\|_2^2 \to \|\mathbf{y}\|_2^2$.
- Feature selection methods attempt to set many coordinates in β to 0. Ridge regularizations encourages coordinates to be small.



POLYNOMIAL EXAMPLES

Fit degree 20 polynomial with varying leves of regularization.



20

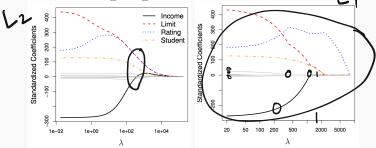
RIDGE REGULARIZATION

How do we minimize:
$$L_{R}(\beta) = (X\beta - y\|_{2}^{2}) + \lambda \|\beta\|_{2}^{2}$$
?
How to solve with $\|X\beta - y\|_{2}^{2} + \lambda \|S\|_{2}^{2}$
 $\nabla L_{R}(\beta) = \chi T (X\beta - j) + \chi \lambda \beta = 0$
 $\chi^{T} X \beta - \chi^{T} y + \lambda \beta = 0$
 $\chi^{T} X \beta + \lambda \beta = \chi^{T} y$
 $(\chi^{T} X + \lambda I) \beta \cdot \chi^{T} y$
 $\beta = (\chi^{T} X + \lambda I)^{-1} \chi^{T} y$
 21

LASSO REGULARIZATION



- As $\lambda \to \infty$, we expect $\|\beta\|_1 \to 0$ and $\|\mathbf{X}\beta \mathbf{y}\|_2^2 \to \|\mathbf{y}\|_2^2$.
- Typically encourages subset of β_i's to go to zero, in contrast to ridge regularization.



Ridge regularization

Lasso Regularization

LASSO REGULARIZATION

Pros:

- Simpler, more interpretable model.
- More intuitive reduction in model order.

Cons:

- No closed form solution because $\|\beta\|_1$ is not differentiable.
- Can be solved with iterative methods, but generally not as quickly as ridge regression.

REGULARIZATION

Notes:

Return at 12:50 from breck.

- Model selection/cross validation used to choose optimal scaling λ on $\lambda \|\beta\|_2^2$ or $\lambda \|\beta\|_1$.
- Often grid search for best parameters is performed in "log space". E.g. consider $[\lambda_1, \ldots, \lambda_q] = 1.5^{[-4, -3, -2, -1, -0, 1, 2, 3, 4]}$.

Regularization methods are <u>not invariant</u> to data scaling. Typically when using regularization we mean center and scale columns to have unit variance. THE BAYESIAN/PROBABILISTIC MODELING PERSPECTIVE

CLASSIFICATION SETUP

Norve Boyes

- Data Examples: $\underline{\mathbf{x}}_1, \dots, \underline{\mathbf{x}}_n \in \mathbb{R}^d$
- Target: $y_1, \ldots, y_n \in \{0, 2, \ldots, q-1\}$ when there are q classes.
 - Binary Classification: q = 2, so each $y_i \in \{0, 1\}$.
 - Multi-class Classification: q > 2.¹

¹Note that there is also multi-label lassification where each data example may belong to more than one class.

- Medical diagnosis from MRI: 2 classes.
- MNIST digits: 10 classes.
- Full Optical Character Regonition: 100s of classes.
- ImageNet challenge: 21,000 classes.

Running example today: Email Spam Classification.

Classification can (and often is) solved using the same **(loss-minimization framework)** we saw for regression.

We won't see that today! We're going to use classification as a window into another way of thinking about machine learning.

Will give new an interesting justifications for tools like <u>regularization</u>. Will also give us an approach for generative ML.

Rest of Today: ML from a Probabilistic Modeling Bayesian Perspective. In a <u>Bayesian</u> or <u>Probabilistic</u> approach to machine learning we always start by conjecturing a

probabilistic model

that plausibly could have generated our data.

- The model guides how we make predictions.
- The model typically has unknown parameters $\vec{\theta}$ and we try to find the most reasonable parameters based on observed data (more on this later in lecture).

Typically we try to keep things simple!



Exercise: Come up with a probabilistic model for the following data set $(x_1, y_1), \ldots, (x_n, y_n)$

 For n NYC apartments: each <u>x</u>_i is the size of the apartment in square feet. Each y_i is the monthly rent in dollars.

What are the unknown parameters of your model. What would be a guess for their values? How would you confirm or refine this guess using data? Dataset: $(x_1, y_1), ..., (x_n, y_n)$

Description: For *n* **NYC apartments**: each x_i is the size of the apartment in square feet. Each y_i is the monthly rent in dollars.

Model:
$$X \sim Uuif(L, U)$$

 $M \sim N(M, 6)$
 $y = C \times + M$
 $fixed$ model

$$Y \in [300, 10000] 58-f4.$$

 $Y := C : + N(\mu, G)$

Dataset: $(x_1, y_1), \ldots, (x_n, y_n)$

Description: For *n* **students**: each

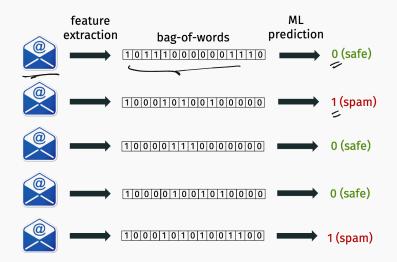
 $x_i \in \{Fresh., Soph., Jun., Sen.\}$ indicating class year. Each $y_i \in \{0, 1\}$ with zero indicating the student has not taken machine learning, one indicating they have.

Model:

Goal:

- Build a probabilistic model for a binary classification problem.
- Estimate parameters of the model.
- From the model derive a classification rule for future predictions (the Naive Bayes Classifier).

SPAM PREDICTION



Both target labels and data vectors are binary.

EMAIL MODEL

Let's create a model that generates spam and non-spam emails. **Observation:** Since b<u>ag-of-w</u>ords <u>features</u> don't care about word order, our model does not need to either.

• Common approach. Assign a probability $p_i \in [0, 1]$ to word *i*. Set $\underline{\mathbf{x}_i = 1}$ with probability p_i , $\underline{\mathbf{x}_i = 0}$ with probability $1 - p_i$. 3° 7_{0} K

$$p_{\text{the}} = . \mathbf{q}$$
 $p_{\text{calendar}} = . \mathcal{V}$ $p_{\text{toothbrush}} = .6665$

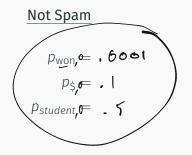
EMAIL MODEL

77,000

Q toothbrush	X 🛱 💽 Active 🔻 🕐 🔅	•
Mail Messages Spaces	From • Any time • Has attachment To • Exclude	Promc > 3
□ • C :	-39 of 39 < >	≡ -
🗆 🕁 Aetna	Inbox Wow, Christophe, have you seen your benefits lately? Electric	10/8/22
🗌 👷 Musco, Jenna	Inbox FW: Cargo e-bikes - , a toothbrush in 15 minutes," said Jolley. "W	5/31/22
🗌 📩 Amazon.com	Amazon Your Amazon.com order #114-7471064-3087458 - Amazon.com	2/14/21
🔲 🖕 Christopher Jenna 3	Inbox first 2 week baby thoughts - keep our toothbrushes in the cabine	12/29/20
🗌 👷 NYU Office of Publi.	Inbox NYU Responds - View in Browser	4/9/20
🗌 ☆ Walgreens.com	10X everyday points = all treat, no trick! - To ensure delivery to your inbo	10/16/19
🗌 🕁 Real Simple	Wayfair's Having a HUGE Sale Right Now—Up to 70% Off	3/17/19
🗌 🛧 Amazon.com	Amazon Your Amazon.com order - Amazon.com Order Confirmation ww	2/3/19
🗌 🕁 Amazon.com	What's new with Alexa? - change my toothbrush in three months." - "Ale	1/12/19
🗌 🛧 Amazon.com	Amazon Your Amazon.com order has shipped (#113-2545543-7095419) - ,	1/2/19
🗌 ☆ Debbie Goodwin	[csail-related] Holiday Giving to begin next week - Toiletries (toothbrush	11/21/18
🗌 🕁 Real Simple	Our Favorite Cooling Bed Sheets Are \$20 Off for Real Simple Readers - $\mbox{-R}$	9/23/18
🔲 👷 erica, Dani., Char. 3	Inbox Last Reminder re: routes - anything (toothbrush, sunblock etc.) w	8/21/18

How can we make this model richer when we take spam into account?

• Different words tend to be more or less frequent in spam or regular emails.





P_{won},⊊ • Y P\$, ⊊ • 3 Pstudent, ₹ • 05

PROBABILISTIC MODEL FOR EMAIL

Probabilistic model for (bag-of-words, label) pair (x, y):

- Set y = 0 with probability \underline{p}_0 , y = 1 with probability $\underline{p}_1 = 1 p_0$.
 - p_0 is probability an email is not spam (e.g. 99%).
 - p₁ is probability an email is spam (e.g. 1%).
- If $\underline{y} = \underline{0}$, for each *i*, set $x_i = 1$ with prob. p_{i_0} .
- If y = 1, for each *i*, set $x_i = 1$ with prob. \underline{p}_{i1} .

Unknown model parameters:

- *p*₀, *p*₁,
- $p_{10}, p_{20}, \dots p_{d0}$, one for each of the *d* vocabulary words.
- $p_{11}, p_{21}, \ldots p_{d1}$, one for each of the *d* vocabulary words.

How would you estimate these parameters?

Reasonable way to set parameters:

- Set p_0 and p_1 to the empirical fraction of not spam/spam emails.
- For each word *i*, set *p*_{i0} to the empirical probability word *i* appears in a <u>non-spam</u> email.
- For each word *i*, set p_{i1} to the empirical probability word *i* appears in a <u>spam</u> email.

Estimating these parameters from previous data examples is the only "training" we will do.

DONE WITH MODELING ON TO PREDICTION

PROBABILITY REVIEW

- **Probability**: p(x) the probability even<u>t x</u> happens.
- Joint probability: p(x,y) the probability that event x and event y happen.
- Conditional Probability p(x | y) the probability x happens given that y happens.

$$\overline{b(x|\lambda)} = b(x, 1)$$

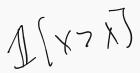
$$\overline{b(x|\lambda)} = b(x, 2)$$

$$\overline{b(x|\lambda)} = b(x, 2)$$

BAYES THEOREM/RULE

$$p(x|y) = \frac{p(y|x)p(x)}{p(y)}$$

Proof:



Given unlabeled input (w, ____), choose the label $y \in \{0, 1\}$ which is <u>most likely</u> given the data. Recall $\mathbf{w} = [0, 0, 1, ..., 1, 0]$.

Classification rule: maximum a posterior (MAP) estimate.

Step 1. Compute:

- $p(y = 0 | \mathbf{w})$: prob. y = 0 given observed data vector \mathbf{w} .
- $p(y = 1 | \mathbf{w})$: prob. y = 1 given observed data vector \mathbf{w} .

Step 2. Output: 0 or 1 depending on which probability is larger.

 $p(y = 0 | \mathbf{w})$ and $p(y = 1 | \mathbf{w})$ are called **posterior** probabilities.

How to compute the posterior? Bayes rule!

$$p(y = 0 | \mathbf{w}) = \frac{p(\mathbf{w} | y = 0)p(y = 0)}{p(\mathbf{w})}$$
(1)

posterior =
$$\frac{\text{likelihood} \times \text{prior}}{\text{evidence}}$$
 (2)

- **Prior:** Probability in class 0 prior to seeing any data.
- **Posterior:** Probability in class 0 <u>after</u> seeing the data.

Goal is to determine which is larger:

$$p(y = 0 | w) = \frac{p(w | y = 0)p(y = 0)}{p(w)}$$
 vs.
$$p(y = 1 | w) = \frac{p(w | y = 1)p(y = 1)}{p(w)}$$

- We can ignore the evidence p(w) since it is the same for both sides!
- p(y = 0) and p(y = 1) already known (computed from training data). These are our computed parameters p_0 , p_1 .

•
$$p(\mathbf{w} \mid y = 0) = ? p(\mathbf{w} \mid y = 1) = ?$$

Consider the example **w** = [0, 1, 1, 0, 0, 0, 1, 0].

Recall that, under our model, index *i* is 1 with probability p_{i0} if we are not spam, and 1 with probability p_{i1} if we are spam.

$$p(\mathbf{w} \mid y = 0) =$$

$$p(w | y = 1) =$$

Final Naive Bayes Classifier

Training/Modeling: Use existing data to compute:

- $p_0 = p(y = 0), p_1 = p(y = 1)$
- For all *i* compute:

•
$$p_{i0} = p(x_i = 1 | y = 0)$$
 and $(1 - p_{i0}) = p(x_i = 0 | y = 0)$

•
$$p_{i1} = p(x_i = 1 | y = 1)$$
 and $(1 - p_{i1}) = p(x_i = 0 | y = 1)$

Prediction:

- For new input **w**:
 - Compute $p(\mathbf{w} \mid y = 0) = \prod_i p(w_i \mid y = 0)$
 - Compute $p(\mathbf{w} \mid y = 1) = \prod_i p(w_i \mid y = 1)$
- Return

arg max
$$[p(\mathbf{w} | y = 0) \cdot p(y = 0), p(\mathbf{w} | y = 1) \cdot p(y = 1)]$$
.

OTHER APPLICATIONS OF THE BAYESIAN PERSPECTIVE The Bayesian view offers an interesting alternative perspective on <u>many</u> machine learning techniques.

Example: Linear Regression.

Probabilistic model:

$$\mathbf{y} = \langle \mathbf{x}, \boldsymbol{\beta} \rangle + \eta$$

where the η drawn from $N(0, \sigma^2)$ is random Gaussian noise.



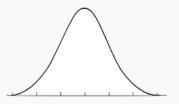
$$Pr(\eta = z) \sim$$

The symbol \sim means "is proportional to".

GAUSSIAN DISTRIBUTION REFRESHER

Names for same thing: Normal distribution, Gaussian distribution, bell curve.

Parameterized by mean μ and variance σ^2 .



 η is a continuous random variable, so it has a <u>probability</u> density function $p(\eta)$ with $\int_{-\infty}^{\infty} p(\eta) d\eta = 1$

$$p(\eta) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}(\frac{\eta-\mu}{\sigma})^2}$$

The important thing to remember is that the PDF falls off exponentially as we move further from the mean.



The normalizing constant in front 1/2, etc. don't matter so much.

Example: Linear Regression. Probabilistic model:

 $y = \langle \mathbf{X}, \boldsymbol{\beta} \rangle + \eta$

where the η drawn from $N(0, \sigma^2)$ is **random Gaussian noise**. The noise is <u>independent</u> for different inputs $\mathbf{x}_1, \dots, \mathbf{x}_n$.

How should be select β for our model?

Also use a Bayesian approach!

Choose $oldsymbol{eta}$ to maximize:

posterior =
$$\Pr(\beta \mid X, y) = \frac{\Pr(X, y \mid \beta) \Pr(\beta)}{\Pr(X, y)} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}$$

In this case, we don't have a prior – no values of $oldsymbol{eta}$ are inherently more likely than others.

Choose $oldsymbol{eta}$ to maximize just the likelihood:

$$\frac{\Pr(X, y \mid \beta)\Pr(\beta)}{\Pr(X, y)} = \frac{\text{likelihood} \times \frac{\text{prior}}{\text{evidence}}$$

This is called the **maximum likelihood estimate**.

Often we think of **X** as fixed and deterministic, and only **y** is generated at random in the model. This is called the <u>fixed</u> <u>design</u> setting. Can also consider a <u>randomized design</u> setting, but it is slightly more complicated.

In the fixed design setting our task of maximizing $\Pr(X, y \mid \beta)$ simplifies to maximizing

 $\max_{\beta} \Pr(\mathsf{y} \mid \beta)$

MAXIMUM LIKELIHOOD ESTIMATE

Data:

$$\mathbf{X} = \begin{bmatrix} - & \mathbf{x}_1 & - \\ - & \mathbf{x}_2 & - \\ \vdots & \\ - & \mathbf{x}_n & - \end{bmatrix} \qquad \qquad \mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

Model: $y_i = \langle \mathbf{x}_i, \boldsymbol{\beta} \rangle + \eta_i$ where $p(\eta_i = z) \sim e^{-z^2/2\sigma^2}$ and η_1, \dots, η_n are independent.

 $\mathsf{Pr}(\mathsf{y} \mid \boldsymbol{\beta}) \sim$

Easier to work with the log likelihood:

$$\arg \max_{\beta} \Pr(\mathbf{X}, \mathbf{y} \mid \beta) = \arg \max_{\beta} \prod_{i=1}^{n} e^{-(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2}$$
$$= \arg \max_{\beta} \log \left(\prod_{i=1}^{n} e^{-(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2} \right)$$
$$= \arg \max_{\beta} \sum_{i=1}^{n} -(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2$$
$$= \arg \min_{\beta} \sum_{i=1}^{n} (y_i - \langle \mathbf{x}_i, \beta \rangle)^2.$$

Conclusion: Choose $\boldsymbol{\beta}$ to minimize:

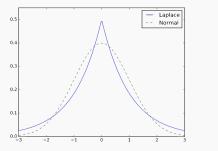
$$\sum_{i=1}^{n} (\mathbf{y}_i - \langle \mathbf{x}_i, \boldsymbol{\beta} \rangle)^2 = \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2.$$

This is a completely different justification for squared loss!

Minimizing the ℓ_2 loss is optimal in a certain sense when you assume your data follows a linear model with i.i.d. Gaussian noise.

BAYESIAN REGRESSION

If we had modeled our noise η as Laplace noise, we would have found that minimizing $\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_1$ was optimal.



$$Pr(\eta = z) \sim$$

Laplace noise has "heavier tails", meaning that it results in more outliers.

This is a completely different justification for ℓ_1 loss.

We can add another layer of probabilistic modeling by also assuming β is random and comes from some distribution, which encodes our <u>prior</u> belief on what the parameters are.

Return to Maximum a posteriori (MAP estimation):

$$\Pr(\beta \mid X, y) = \frac{\Pr(X, y \mid \beta) \Pr(\beta)}{\Pr(X, y)}$$

Assume values in $\beta = [\beta_1, \dots, \beta_d]$ come from some distribution.

- **Common model:** Each β_i drawn from $N(0, \gamma^2)$, i.e. normally distributed, independent.
- Encodes a belief that we are unlikely to see models with very large coefficients.

Goal: choose β to maximize:

$$\Pr(\beta \mid X, y) = \frac{\Pr(X, y \mid \beta) \Pr(\beta)}{\Pr(X, y)}.$$

- We can still ignore the "evidence" term Pr(X, y) since it is a constant that does not depend on β.
- $\Pr(\beta) = \Pr(\beta_1) \cdot \Pr(\beta_2) \cdot \ldots \cdot \Pr(\beta_d)$
- · $\Pr(eta) \sim$

BAYESIAN REGULARIZATION

Easier to work with the log likelihood:

$$\arg \max_{\beta} \Pr(\mathbf{X}, \mathbf{y} \mid \beta) \cdot \Pr(\beta)$$

$$= \arg \max_{\beta} \prod_{i=1}^{n} e^{-(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2} \cdot \prod_{i=1}^{n} e^{-(\beta_i)^2 / 2\gamma^2}$$

$$= \arg \max_{\beta} \sum_{i=1}^{n} -(y_i - \langle \mathbf{x}_i, \beta \rangle)^2 / 2\sigma^2 + \sum_{i=1}^{d} -(\beta_i)^2 / 2\gamma^2$$

$$= \arg \min_{\beta} \sum_{i=1}^{n} (y_i - \langle \mathbf{x}_i, \beta \rangle)^2 + \frac{\sigma^2}{\gamma^2} \sum_{i=1}^{d} (\beta_i)^2 / \sigma^2.$$

Choose $\boldsymbol{\beta}$ to minimize $\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \frac{\sigma^2}{\gamma^2} \|\boldsymbol{\beta}\|_2^2$.

Completely different justification for ridge regularization!

Test your intuition: What modeling assumption justifies LASSO regularization: min $\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$?